Novel Statistical Methods for Data Harmonization

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March 2018

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Where innovation starts

Contents

- IPD Meta-analysis
- Harmonization approaches
- Standardization methods
- Simulation study
 - Data generation
 - Analysis method
 - Results
 - Conclusions



- The gold standard in meta-analysis is "individual participant data (IPD)" meta-analysis
- Advantages of IPD
 - Correcting for confounders at individual level
 - Testing interactions of covariates with exposure is more powerful at individual level
 - Subgroup analysis to investigate heterogeneity or consistency
- There are two analysis forms of IPD:
 - Two-stage IPD
 - One-stage IPD



Two-stage Analysis:

- Similar to an analysis of aggregate data (AD) meta-analysis
- Same statistical model is fitted to each study
- Each study fits the same covariates
 - Estimates can be corrected when not all covariates are available (Fibrinogen Study, 2009)
- Parameter estimates are combined via regular meta-analysis techniques
 - Weighted averages (e.g. DerSimonian-Laird)
 - Maximum likelihood estimates

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One-stage Analysis:

- Is one statistical analysis of all data addressing random and fixed effects
 - Studies are considered random
- Requires more sophisticated statistical tools and may be numerically more challenging
 - Analysis may require several hierarchical structures of random effects (especially for longitudinal data)
 - In some cases the analysis must be executed in a *federated* or *distributed* way (data can not be pooled into one location)

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Federated Meta-Analysis:

- Expectation-maximization algorithm
 - Step 0: Choose starting values of all parameters
 - Step 1: E-Step: use the estimates from the previous step to estimate random effects
 M-Step: Using the result from the E-step determine fixed effects parameters
 - Evaluate: how much the estimates has changed
 - If the changes are small enough \rightarrow convergence
 - If the changes are still to large → conduct step 1 using the last available estimates
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 - EM uses study summary statistics

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Federated Meta-Analysis: Example

- Response: Systolic blood pressure
- Exposure: Noise
- Confounders: Age, Sex, PM10 (particulate matter)
- Two cohorts: HUNT and LifeLines

	β ₀	β_{AGE}	β_{SEX}	β _{PM10}	β_{NOISE}	τ ²	σ^2
EM-0	1	1	1	1	1	1	1
EM-Final	111.59	0.4141	-7.255	0.04627	-0.01351	1.8992	217.43
EM-0	1	1	1	1	1	0	1
EM-Final	114.68	0.4143	-7.2473	-0.16617	-0.00300	0	217.45

 EM is dependent on starting value and it makes a difference in estimation

- One-stage IPD requires more effort since similar variables can not just simply be pooled
- Variables from different studies may not contain the exact same information
- Example on memory:

Study	HUI	RAVLT	Free BRCP	Cued BRCP
CCHS	5	15	-	-
CSHA	-	15	12	12
NuAge	-	-	16	16

- RAVLT=Rey auditory verbal learning test
- BRCP=Buschke cued recall procedure
- HUI=Health utility score

Correct	CCHS (r	า=7107)	C	CSHA (n=1730))	NuAge (n=432)	
# words	HUI	Rey	Rey	Free B	Cued B	Free B	Cued B
0	71.1	1.17	4.57	4.28	0.23	0.23	0
1	2.04	2.81	7.28	2.31	0.29	0	0
2	20.9	8.68	13.1	3.70	0.58	0.69	0
3	4.94	15.6	21.6	5.09	0.40	3.24	0
4	0.99	20.3	17.7	8.61	0.58	6.48	0.23
5	0.01	10.2	13.1	10.9	0.81	13.4	0
6	Na	15.3	5.32	14.0	1.27	14.1	0.69
7	Na	8.33	2.60	17.4	2.08	16.0	0.93
8	Na	4.54	1.05	14.3	2.66	14.8	0.93
9	Na	1.89	0.06	10.7	4.22	12.3	2.31
10	Na	0.79	0.17	6.42	7.11	8.80	1.16
11	Na	0.27	0	2.02	16.5	4.63	5.56
12	Na	0.11	0	0.29	63.2	3.24	7.64
13	Na	0.03	0	Na	Na	1.85	10.6
14	Na	0	0	Na	Na	0.23	18.8
15	Na	0	0	Na	Na	0	19.4
16	Na	Na	Na	Na	Na	0	31.7

- Bridge variables are variables that make it possible to connect studies
 - RAVLT connects CCHS with CSHA
 - BCRP connects CSHA with NuAge
 - Thus all studies are connected
- Different Harmonization methods
 <u>With bridge variables</u> <u>Withou</u>
 - Sequential calibration
 - Latent variable models
 - Imputation methods

Without bridge variables

- Algorithmic methods
- Standardization or normalization
- Imputation methods

- Algorithmic harmonization:
 - Transform y into categories (low, medium, high)
 - Thresholds are test and demographic specific
- Standardization or normalization:
 - Change the scores to a common scale:
 - Min-max scaling: $(y y_{min})/(y_{max} y_{min})$
 - Z-scores: $(y \overline{y})/s$
 - T-scores: $(y \overline{y})/s$ corrected for covariates
 - C-scores: $(y \bar{y}_c)/s_c$ corrected for covariates, with \bar{y}_c an average of a well-defined control group
 - Quantile normalization: $F^{-1}(y)$

- Calibration on responses
 - Sequential relations are formed: $y_2 = \psi(y_1)$
 - Requires "bridge variables" to connect studies
- Latent variable models:
 - Factor analysis or item response theory models
 - The latent variable is the harmonized variable
 - Requires "bridge variables" to connect studies
- Multiple imputation methods:
 - Must deal with sporadic and systematic missingness
 - Makes all variables available in all studies
 - Requires other bridge variables

Latent variable models

- Let Z_i be the ability or latent variable on subject i
- Let Y_{hi} be the score for test h on subject i
- We assume that Y_{hi} given Z_i is binomial $Y_{hi}|Z_i \sim Bin(N_h, p_h(Z_i))$
- The different tests are harmonizable when there exist a function ψ such that

 $p_1(Z_i) = p_2(\psi(Z_i)) \text{ or } p_2(Z_i) = p_1(\psi(Z_i))$

 Calibration model on latent variable similar to inches and centimeters: 1 inch = 2.54 cm

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• We choose $\psi(z) = a + bz$

Latent variable models

- The distribution of Z_i is considered normal
 - Mean: $\beta_0 + \sum_{m=1}^p \beta_m x_{m,i}$
 - Variance: $\gamma_0 + \sum_{m=1}^p \gamma_m x_{m,i}$
- The probability p_h is $logit(p_h(z)) = \mu_h + \eta_h z$

Variable		CCHS	CSHA	NuAge	Dualua	
variable	Level	(n=7107)	(n=1730)	(n=432)	P-value	
Age [mean(sd)]	Numeric	73.2 (5.85)	79.7 (6.96)	73.7 (3.95)	<0.001	
C and [0/]	Male	42.3%	37.3%	46.3%	<0.001	
Sex [70]	Female	57.7%	62.7%	53.7%		
Education	Low	19.0%	48.6%	15.3%		
Education [%]%	Medium	37.7%	35.8%	39.6%	<0.001	
	High	43.3%	15.6%	45.1%		

Latent variable models:

Daramators	Mean			Variance			
Parameters	CCHS	CSHA	NuAge	CCHS	CSHA	NuAge	
Intercept	0.955	2.653	1.009	-3.124	-2.161	-2.263	
HUI	2.813	NA	NA	Zero (due to weak invariance)			
RAVLT	NA	-1.444	NA				
Cued Buschke	NA 🕻	2.693	2.291				
Med. Education	0.201	0.049	0.160	-0.063	0.106	0.207	
High education	0.356	0.176	0.357	-0.010	0.144	0.184	
Sex	0.210	-0004	0.449	0.148	0.286	0.018	
Age	-0.029	-0033	-0.022	0.023	0.021	0.020	

• Model fit: $\chi^2/df = 1.114$

- Without bridge variables there is no way to check "content equivalence"
 - BMI and IQ can be standardized, normalized or categorized, but they do not become exchangeable
- Bridge variables could check calibration or latent variable invariance principles
 - Consistency of the relation between free and cued BCRP can be verified for CSHA and NuAge
 - Sparsity restricts the verification of all invariances
 - Harmonization invariance is unequal to factorial measurement invariance but it is similar to differential item functioning

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- Compare standardization methods
 - T-score (corrected for confounders)
 - C-score (corrected for confounders)
- Harmonization setting
 - Response: Memory construct
 - Exposure: physical activity (low, medium, high)
 - Confounders: Age, gender, and education
- Two stage IPD meta-analysis
 - Pooled effect size: Hedges g (adjusted and unadjusted for confounders)
 - Measure of heterogeneity: I^2

Calculations of T- and C-score

- Let Y_i be memory scale for participant *i* in one study
- T-score

$$Z_{i} = 10 + 3(Y_{i} - \overline{Y})/S$$

$$Z_{i} = \alpha_{0} + \sum_{k=1}^{p} \alpha_{k} x_{ik} + e_{i}, \quad e_{i} \sim N(0, \sigma_{e}^{2})$$

$$T_{i} = 50 + 10(Z_{i} - \hat{Z}_{i})/\hat{\sigma}_{e}$$

- With $\overline{Y} = \sum_{i=1}^{n} Y_i$, $S = \sum_{i=1}^{n} (Y_i \overline{Y})^2 / (n-1)$, \hat{Z}_i the predicted Z_i , and $\hat{\sigma}_e$ an estimate of σ_e
- C-score: $C_i = (Y_i \overline{Y}_C)/S_C$
 - With \overline{Y}_{C} and S_{C} the average and standard deviation of 70 to 74 years old and high $TU e^{\frac{Technische Universiteit}{Eindhoven}}$ educated females 30-3-2018 PAGE 17

Calculations of hedges g

- Unadjusted effect size of PA on memory per study
 - Three levels: Low-Medium, Low-High, Medium-High
- Let (w_k, τ_k, n_k) be the mean, standard deviation, and sample size for T- or C-score at PA level k
- Effect size for PA level k versus l:
 - Mean difference: $d_{kl} = w_k w_l$
 - Pooled standard deviation: $s_{kl}^2 = \frac{(n_k 1)\tau_k^2 + (n_l 1)\tau_l^2}{n_k + n_l 2}$

• Hedges g:
$$g_{kl} = \left(1 - \frac{3}{4(n_k + n_l) - 9}\right) \frac{d_k}{s_{kl}}$$

• Variance: $v_{kl}^2 = \frac{n_k + n_l}{n_k n_l} + \frac{g_{kl}^2}{2(n_k + n_l)}$

Calculations of hedges g

- Analysis of T- or C-score corrected for age, sex, education uses linear regression per study
 - Residual variance changes with PA levels and replaces τ_k^2 and τ_k^2
 - Least square means of PA in regression analysis are used as substitutes for difference d_{kl}
 - Samples sizes n_k and n_l remain the same
- DerSimonian & Laird to combine hedges g's
- Heterogeneity: $I_{kl}^2 = (Q_{kl} m + 1)/Q_{kl}$

 $Q_{kl} = \sum_{r=1}^{m} v_{kl,r}^{-2} (d_{kl,r} - \bar{d}_{kl})^2$

- Age, gender, and education were independently drawn from normal and Bernoulli distribution
- Physical activity was simulated from logistic distribution with mean dependent on confounders
 - Then physical activity was set to three categories
- Latent memory followed a normal distribution
 - mean depend on physical activity and confounders
 - Variance depend on confounders
- Memory score given latent variable followed a binomial distribution with logistic link function
- Date generation was done per study TU e Eindhoven University

- Age $A \sim N(\mu, \sigma^2)$
- Gender $S \sim B(\pi_S)$
- Education $E \sim U(0,1)$
 - Low education: $E_L = 1$: 0
 - Medium education: E
 - High education:

 $E_L = 1: \qquad 0 \le U \le \pi_L$ $E_M = 1: \qquad \pi_L < U \le \pi_M$ $E_H = 1: \qquad \pi_M < U \le 1$

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Parameters	Homogeneous populations			Heterogeneous populations		
	Study 1	Study 2	Study 3	Study 1	Study 2	Study 3
μ_A	75	75	75	70	80	75
σ_A	6	6	6	6	7	4
π_s	0.40	0.40	0.40	0.40	0.35	0.45
π_L	0.30	0.30	0.30	0.20	0.50	0.15
π_M	0.70	0.70	0.70	0.60	0.85	0.55
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• Physical Activity $X \sim Logistic(\mu_{PA}, 1)$

 $\mu_{PA} = -0.03A - 0.5S - 0.6E_L - 0.3E_M$

- The influence of confounder on physical activity was the same in each study
- Physical activity was categorized into three categories:
 - Low PA: $P_L = 1$: $X \le 1.0$
 - Medium PA: $P_M = 1$: $1.0 < X \le 3.5$
 - High PA: $P_H = 1$: 3.5 < X



• Latent Memory $Z \sim N(\mu_M, \sigma_M^2)$

 $\mu_M = \beta_0 + \beta_A A + \beta_S S + \beta_L E_L + \beta_M E_M + \beta_1 P_L + \beta_2 P_M$

 $\log \sigma_M = \gamma_0 + \gamma_A A + \gamma_S S + \gamma_L E_L + \gamma_M E_M$

- Memory scale $Y \sim Bin(N, p(Z))$
 - Number of words tested N
 - Study 1: N = 15
 - Study 2: N = 12
 - Study 3: N = 16
 - Logistic function for probability p(Z) = exp(Z)/(1 + exp(Z))



Parameters of confounders for memory

Daramatara	Homogeneous effects			Heterogeneous effects			
Parameters	Study 1	Study 2	Study 3	Study 1	Study 2	Study 3	
β _o	2.5	2.5	2.5	1.5	2.5	3.5	
β_A	-0.025	-0.025	-0.025	-0.025	-0.035	-0.020	
β_s	0.20	0.20	0.20	0.20	0	0.45	
β_L	-0.25	-0.25	-0.25	-0.35	-0.15	-0.35	
β_M	-0.15	-0.15	-0.15	-0.15	-0.10	-0.20	
γ ₀	-2.0	-2.0	-2.0	-3.0	-2.0	-2.0	
γ_A	0.02	0.02	0.02	0.02	0.02	0.02	
γ_S	0.15	0.15	0.15	0.15	0.25	0	
γ_L	-0.15	-0.15	-0.15	0	-0.15	-0.20	
Υ _M	-0.05	-0.05	-0.05	0	-0.05	0	

Parameters of physical activity on memory

Parameters	Homogeneous effects			Heterogeneous effects		
	Study 1	Study 2	Study 3	Study 1	Study 2	Study 3
β_1	-0.50	-0.50	-0.50	-0.60	-0.30	-0.10
β_2	-0.25	-0.25	-0.25	-0.30	0	-0.10

Simulation Study Results

Unadjusted effect sizes

Effect of PA	Population	Confounder Effect on Memory	Type of Outcome	Effect size	Power	I ²
Homo	Homo	Homo	T-score	0.57	100	16.6
Homo	Homo	Homo	C-score	0.53	100	17.1
Homo	Homo	Hetero	T-score	0.62	100	87.6
Homo	Homo	Hetero	C-score	0.57	100	82.7
Homo	Hetero	Homo	T-score	0.58	100	27.7
Homo	Hetero	Homo	C-score	0.54	100	30.0
Homo	Hetero	Hetero	T-score	0.62	100	91.9
Homo	Hetero	Hetero	C-score	0.57	100	89.8
Hetero	Homo	Homo	T-score	0.39	73.2	95.4
Hetero	Homo	Homo	C-score	0.36	56.2	95.6
Hetero	Homo	Hetero	T-score	0.46	19.9	98.2
Hetero	Homo	Hetero	C-score	0.41	13.5	98.0
Hetero	Hetero	Homo	T-score	0.40	62.8	96.0
Hetero	Hetero	Homo	C-score	0.37	46.9	96.0
Hetero	Hetero	Hetero	T-score	0.47	10.2	98.4
Hetero	Hetero	Hetero	C-score	0.42	7.0	98.3

Simulation Study Results

Adjusted effect sizes

Effect of PA	Population	Confounder Effect on Memory	Type of Outcome	Effect size	Power	I ²
Homo	Homo	Homo	T-score	0.59	100	18.4
Homo	Homo	Homo	C-score	0.59	100	18.4
Homo	Homo	Hetero	T-score	0.65	100	88.5
Homo	Homo	Hetero	C-score	0.65	100	88.5
Homo	Hetero	Homo	T-score	0.60	100	29.9
Homo	Hetero	Homo	C-score	0.60	100	29.9
Homo	Hetero	Hetero	T-score	0.64	100	92.5
Homo	Hetero	Hetero	C-score	0.64	100	92.5
Hetero	Homo	Homo	T-score	0.41	72.7	95.8
Hetero	Homo	Homo	C-score	0.41	72.7	95.8
Hetero	Homo	Hetero	T-score	0.48	19.4	99.4
Hetero	Homo	Hetero	C-score	0.48	19.4	99.4
Hetero	Hetero	Homo	T-score	0.41	61.5	96.4
Hetero	Hetero	Homo	C-score	0.41	61.5	96.4
Hetero	Hetero	Hetero	T-score	0.48	9.2	98.6
Hetero	Hetero	Hetero	C-score	0.48	9.2	98.6

Simulation Study Conclusions

- Pooled effects of PA on memory for T and C are
 - Different when scores are unadjusted
 - Almost identical for adjusted effect sizes
 - A bit affected by heterogeneity in confounders
- Unadjusted effect sizes show less study heterogeneity for C-scores than for T-scores
- Heterogeneity in homogeneous effect sizes of PA on memory with T- and C-scores
 - Is disturbed by heterogeneous effect sizes of the confounders on memory (adjusted and unadjusted)
 - Is only little affected by heterogeneityTU e Technische Universiteit Eindhoven University of Technology
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Simulation Study Conclusions



Biometrics & Biostatistics International Journal

Statistical Harmonization Methods in Individual Participants Data Meta-Analysis are Highly Needed

Editorial

Meta-analysis has a long history within the medical sciences and epidemiology [1-3]. The main goal of a meta-analysis is to improve the precision of a specific effect size of a treatment or an exposure on a clinical or disease outcome by pooling or combining multiple studies. It is frequently conducted within a systematic review of the scientific literature to guarantee that studies with appropriate information are not ignored or overlooked and to make sure that the pooled estimate represents an unbiased and precise estimate of the true effect size. Moreover, the pooled effect estimate is evaluated in the context of study heterogeneity. In case substantial heterogeneity in the effect sizes across studies is present, the pooled estimate is considered less reliable or even questionable. Thus not only the pooled estimate must be unbiased, also the estimate of the measure

Editorial

Volume 3 Issue 3 - 2016

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Received: January 28, 2016 | Published: February 01, 2016